



Effect of Inspection Error on CUSUM Chart for Truncated Negative Binomial Distribution

U. Mishra ^{a*} and J. R. Singh ^b

^a Shri Vaishnav Vidyapeeth Vishwavidyalaya, Indore (M.P.), India.

^b School of Studies in Statistics, Vikram University, Ujjain (M.P.), India.

Authors' contributions

This work was carried out in collaboration between both authors. Both authors read and approved the final manuscript.

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Abstract

In this paper, a mathematical investigation has been done for Cumulative Sum control Chart (CSCC) of Truncated Negative Binomial Distribution (TNBD) under inspection error using the sequential probability ratio test (SPRT). Average Run Length (ARL), d (lead distance) and ϕ (angle of the mask) is calculated for different values of error rates and different values of parameter of the distribution.

Keywords: TNBD; CSCC; ARL; inspection error; SPRT.

1 Introduction

In statistics, truncated distributions are conditional distributions that effect from confining the area of additional probability distribution. Truncated distributions occur in realistic statistics in cases where the capability to record or even to know about, occurrences is restricted to values which lie within a specified range. The Poisson is often the standard distribution regarded as modeling random counts. As such, for analyzing control limits for processes with discrete responses, usual procedures are based on the Poisson distribution. Poisson distribution has the distinguished property that the mean and variance of the distribution is equal. Though, for several processes, the Poisson distribution offers an unsatisfactory model. Different kinds of procedures can make distributions of counts which are not effectively modeled by the Poisson distribution. Such

*Corresponding author: Email: uttamamishra@gmail.com;

processes comprise conditions where the intensity rates of the counts differ at random over time. Negative binomial distribution is a typical expansion of the Poisson distribution and authorizes for over-dispersion relative to the Poisson. The negative binomial distribution can be resultant of numerous models. The negative binomial is also resulting as a mixture of Poisson distributions however, usually resulting as a generalization of the geometric distribution. Applications of the negative binomial distribution are extensive. The negative binomial distribution is extensively used for the description of data excessively assorted to be fitted by a Poisson distribution. However, observed samples may be truncated, in the sense that the number of individuals falling into the zero class cannot be determined. Samford [1] have illustrated the use of truncated negative binomial distribution over Poisson distribution by considering the method of moments. Hoffman [2] considered the control limits of negative binomial distribution for count data with discrete responses based on Poisson distribution. Chakraborty et al. [3] investigated the effect of measurement error on the power and ARL of control chart for ZTNBD based on standardized normal variates.

The CUSUM chart is exercises to examine the mean of a process based on samples taken from the process at given times. The measurements of the samples in a given time comprises a subgroup rather examining the mean of every subgroup independently, the CUSUM chart illustrates the accumulated information of existing and earlier samples. This is the reason why CUSUM chart is usually better than the \bar{X} –chart for detecting small shifts in the mean of a process. The CUSUM charts rely on the requirement of a target value and a known or consistent estimate of the standard deviation. This is the reason; the CUSUM chart is better used past the process control has been recognized. The CUSUM chart usually signals an out-of-control process by a growing or sliding drift of the cumulative sum until it crosses the boundary. An assignable cause is suspected whenever the CUSUM chart indicates an out-of-control process. Hoffman [2] calculated exact and approximate control limits for count data based on the negative binomial distribution. Sankle et al. [4] presented CSCC for truncated normal distribution considering the effect of inspection error. Sayyed and Singh [5] considered CSCC for binomial parameters under the effect of inspection error where the underlying distribution is Poisson. Singh and Mishra [6] have considered the effect of inspection error on singly truncated Binomial distribution. Chakraborty and khurshid [7] studied the connection between apparent fraction defective (AFD) and true fraction defective (TFD) on the power of control chart.

The point of inspection and quality control in manufacturing operations is to avoid manufacturing mistakes. Without regular manufacturing error inspection, many mistakes would get through that could have catastrophic consequences for a business. Ya-Hui Lin et al. [8] develops an integrated model of production lot-sizing, maintenance and quality for considering the possibilities of inspection errors, preventive maintenance (PM) errors and minimal repairs for an imperfect production system with increasing hazard rates. Sarkar et al. [9] considered that product inspection performs at any arbitrary time of the production cycle and after the inspection, all defective products produced until the end of the production run are fully reworked. Due to some misclassification during inspection, from the inspector’s side two types of inspection errors as Type I and Type II are considered to make the model more realistic rather than existing models.

In this paper, we have seen the effect of inspection error on CSCC for TNBD. ARL, d , and ϕ is calculated for different values of parameter of TNBD under inspection error.

2 Truncated Negative Binomial Distribution (TNBD) under Inspection Error

The negative binomial distribution has the following probability density function (pdf):

$$f(x; r, p) = \begin{cases} \binom{x+r-1}{r-1} p^r q^x & ; x = 0, 1, 2, 3 \dots \\ 0, & \text{otherwise,} \end{cases} \tag{1}$$

where r and p are the parameters of Negative Binomial distribution which satisfy $0 < p < 1$ and $r = 0, 1, 2, 3 \dots$ If $p = \frac{1}{Q}$ and $q = \frac{P}{Q}$ then we have $Q - P = 1$, then the equation (1) will be written as:

$$f(x; r, p) = \begin{cases} \binom{x+r-1}{r-1} \left(\frac{p}{Q}\right)^x \left(1 - \frac{p}{Q}\right)^r & ; x = 0, 1, 2, 3 \dots \\ 0, & \text{otherwise.} \end{cases} \tag{2}$$

To obtain the corresponding probabilities for the truncated distribution the equation (2) must be divided by $(1 - f(x = 0; r, p))$ where $f(x = 0; r, p) = \left(1 - \frac{p}{Q}\right)^r$, it follows that:

$$f(x; r, p) = \begin{cases} \binom{x+r-1}{r-1} \left(\frac{1}{1-Q^{-r}}\right) \left(\frac{p}{Q}\right)^x \left(1 - \frac{p}{Q}\right)^r & ; x = 1, 2, 3 \dots \\ 0, & \text{otherwise.} \end{cases} \tag{3}$$

The mean and variance of the above distribution are as follows:

$$\left. \begin{aligned} \text{Mean} = E(X) &= rP(1 - Q^{-r})^{-1} \\ \text{Variance} = V(X) &= \frac{rPQ}{(1-Q^{-r})} \left[1 - \frac{rP}{Q} ((1 - Q^{-r}) - 1)\right] \end{aligned} \right\} \tag{4}$$

Moreover, inspection of attributes are differentiated by two decision variables in which every item is analyzed and classified as good or faulty. Two types of error are possible, an item which is good but categorized as faulty, e_1 , or an item which is faulty but categorized as good, e_2 . If p is true fraction defective and p' is the apparent fraction defective, then we set:

$$p' = p(1 - e_2) + (1 - p)e_1, \tag{5}$$

with both e_1 and e_2 estimated.

3 CSCC for TNBD under Inspection Error

Let $x_i (i = 1, 2, 3, \dots)$ are the independent random variables with the distribution given by equation (3). The sequential probability ratio test (SPRT) distinguishes between the hypotheses $H_0: P' = P'_0$ and $H_1: P' = P'_1 (> P'_0)$ has the value:

$$\frac{f(x_1, x_2, \dots, x_n | P'_1 r)}{f(x_1, x_2, \dots, x_n | P'_0 r)} = \prod_{i=1}^n \left(\frac{1 - (Q'_0)^{-r}}{1 - (Q'_1)^{-r}}\right)^{-n} (1 - (Q')^{-r})^{-1} \left(\frac{P'_1 Q'_0}{P'_0 Q'_1}\right)^{\sum x_i} \left(\frac{Q'_0(Q'_1 - P'_1)}{Q'_1(Q'_0 - P'_0)}\right)^{nr} \tag{6}$$

The SPRT has the continuation region is then

$$\left(\frac{\beta}{1-\alpha}\right) < \prod_{i=1}^n \left(\frac{1 - (Q'_0)^{-r}}{1 - (Q'_1)^{-r}}\right)^{-n} (1 - (Q')^{-r})^{-1} \left(\frac{P'_1 Q'_0}{P'_0 Q'_1}\right)^{\sum x_i} \left(\frac{Q'_0(Q'_1 - P'_1)}{Q'_1(Q'_0 - P'_0)}\right)^{nr} < \left(\frac{1-\beta}{\alpha}\right), \tag{7}$$

that is,

$$\log \left(\frac{\beta}{1-\alpha}\right) < n \log \left(\frac{1 - (Q'_0)^{-r}}{1 - (Q'_1)^{-r}}\right) + r \log(1 - Q') + \sum_{i=1}^n x_i \log \left(\frac{P'_1 Q'_0}{P'_0 Q'_1}\right) + nr \log \left(\frac{Q'_0(Q'_1 - P'_1)}{Q'_1(Q'_0 - P'_0)}\right) < \log \left(\frac{1-\beta}{\alpha}\right). \tag{8}$$

Now consider the right hand side of the equation (8), we have:

$$n \log \left(\frac{1-(Q'_0)^{-r}}{1-(Q'_1)^{-r}} \right) + r \log(1-Q') + \sum_{i=1}^n x_i \log \left(\frac{P'_1 Q'_0}{P'_0 Q'_1} \right) + nr \log \left(\frac{Q'_0(Q'_1-P'_1)}{Q'_1(Q'_0-P'_0)} \right) < \log \left(\frac{1-\beta}{\alpha} \right), \tag{9}$$

$$\sum_{i=1}^n x_i \log \left(\frac{P'_1 Q'_0}{P'_0 Q'_1} \right) < \log \left(\frac{1-\beta}{\alpha} \right) - n \log \left(\frac{1-(Q'_0)^{-r}}{1-(Q'_1)^{-r}} \right) - nr \log \left(\frac{Q'_0(Q'_1-P'_1)}{Q'_1(Q'_0-P'_0)} \right) - r \log(1-Q'). \tag{10}$$

If we plot points $(n, \sum_{i=1}^n x_i)$ then the boundary line between the continuation region and the acceptance region for H_1 has the equation:

$$\sum_{i=1}^n x_i < \frac{\log \left(\frac{1-\beta}{\alpha} \right) - n \log \left(\frac{1-(Q'_0)^{-r}}{1-(Q'_1)^{-r}} \right) - nr \log \left(\frac{Q'_0(Q'_1-P'_1)}{Q'_1(Q'_0-P'_0)} \right) - r \log(1-Q')}{\log \left(\frac{P'_1 Q'_0}{P'_0 Q'_1} \right)}, \tag{11}$$

or using the approximation of $\beta = 0$, we have:

$$\sum_{i=1}^n x_i < \frac{-\log \alpha - n \log \left(\frac{1-(Q'_0)^{-r}}{1-(Q'_1)^{-r}} \right) - nr \log \left(\frac{Q'_0(Q'_1-P'_1)}{Q'_1(Q'_0-P'_0)} \right) - r \log(1-Q')}{\log(P'_1 Q'_0) - \log(P'_0 Q'_1)}. \tag{12}$$

To construct the CUSUM chart and the V-Mask, the lead distance d and the angle ϕ of the mask is given by:

$$d = \frac{-\log \alpha}{\left[\log \left(\frac{1-(Q'_1)^{-r}}{1-(Q'_0)^{-r}} \right) + r \log \left(\frac{Q'_1(Q'_0-P'_0)}{Q'_0(Q'_1-P'_1)} \right) \right]}, \tag{13}$$

and

$$\phi = \tan^{-1} \left[\frac{\log \left(\frac{1-(Q'_1)^{-r}}{1-(Q'_0)^{-r}} \right) + r \log \left(\frac{Q'_1(Q'_0-P'_0)}{Q'_0(Q'_1-P'_1)} \right)}{\log(P'_1 Q'_0) - \log(P'_0 Q'_1)} \right]. \tag{14}$$

The approximate formula for the average run length (ARL) for detecting a shift for the parameter P from P'_0 to P'_1 for known r is given by,

$$ARL = \frac{-\log \alpha}{\left[\log \left(\frac{1-(Q'_0)^{-r}}{1-(Q'_1)^{-r}} \right) + \left(\frac{rP'_1}{1-(Q'_1)^{-r}} \right) \log \left(\frac{P'_1 Q'_0}{P'_0 Q'_1} \right) + r \log \left(\frac{Q'_0(Q'_1-P'_1)}{Q'_1(Q'_0-P'_0)} \right) \right]}. \tag{15}$$

Values of the d , ϕ and ARL of CUSUM chart are determined for a different combinations of the values of P , r and α for controlling the parameters P when r is known. These values are shown in Table 1.1 to Table 1.6.

4 Illustration and Conclusion

For the purpose of numerical illustration, we have considered five cases: $(e_1, e_2) = (0,0), (0.03,0.3), (0.03,0.1), (0.01, 0.05), (0.005, 0.02)$. Table- 1.1 shows that values of the lead distance d and ARL for different values of α and k in error free case. In Table- 1.2 to Table- 1.5 we have taken different combinations of error rates. It is seen that the value of d and ARL decreases as the value of α and k increases. But if we compare error free case given in Table-1.1 to other error rates shown in Table 1.2 to Table- 1.5, it is seen that for $k=1$ and $\alpha = 0.05$, the value of d and ARL is 4.31 and 17.54 respectively when $(e_1, e_2) = (0,0)$ and for the same value of k and α , the values of $d= 4.41, 4.53, 4.64, 5.66$ and $ARL= 18.22, 19.10, 20.69, 26.55$ when $(e_1, e_2) = (0.005, 0.02), (0.03,0.3), (0.03,0.1), (0.01, 0.05)$ respectively. The above discussion shows that as the error rate increases the value of d and ARL increases, which is a good symptom to detect assignable causes less often in production process.

Table 1.1. Value of d and ARL when $(e_1, e_2) = (0, 0)$

(P'_0, P'_1)		α														
		0.05			0.025			0.01			0.005			0.001		
		<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>
(1, 2)	d →	4.31	3.04	2.27	5.31	3.75	2.80	6.62	4.68	3.50	7.62	5.38	4.02	9.94	7.03	5.25
	ARL →	17.54	9.49	6.20	21.60	11.69	7.64	26.96	14.60	9.53	31.03	16.80	10.97	40.45	21.90	14.30
(1, 3)	d →	2.72	1.86	1.36	3.35	2.29	1.67	4.19	2.86	2.09	4.82	3.29	2.41	6.28	4.29	3.14
	ARL →	5.72	3.04	1.98	7.05	3.74	2.44	10.12	4.67	3.05	10.13	5.37	3.50	13.20	7.01	4.57
(1, 4)	d →	2.16	1.44	1.04	2.66	1.77	1.28	3.32	2.21	1.60	3.82	2.54	1.84	4.98	3.32	2.40
	ARL →	3.11	1.63	1.06	3.83	2.01	1.35	4.77	2.50	1.63	5.49	2.88	1.88	7.16	3.76	2.46
(1,5)	d →	1.85	1.21	0.87	2.28	1.49	1.07	2.85	1.86	1.33	3.28	2.14	1.54	4.28	2.80	2.01
	ARL →	2.04	1.06	0.69	2.51	1.31	0.85	3.13	1.63	1.06	3.60	1.87	1.23	4.70	2.45	1.60

Table 1.2. Value of d and ARL when $(e_1, e_2) = (0.01, 0.05)$

(P'_0, P'_1)		α														
		0.05			0.025			0.01			0.005			0.001		
		<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>
(1, 2)	d →	4.53	3.17	2.35	5.59	3.91	2.90	6.97	4.87	3.62	8.02	5.61	4.17	10.47	7.32	5.43
	ARL →	19.10	10.30	6.72	23.52	12.68	8.28	29.36	15.83	10.33	3.78	18.22	4.88	44.04	23.75	15.50
(1, 3)	d →	2.86	1.93	1.41	3.52	2.38	1.74	4.40	2.98	2.17	5.06	3.42	2.49	6.61	4.47	3.25
	ARL →	6.28	3.32	2.17	7.74	4.09	2.67	9.66	5.11	3.34	11.11	5.88	3.84	14.49	7.67	5.01
(1, 4)	d →	2.27	1.50	1.08	2.79	1.85	1.33	3.49	2.31	1.66	4.01	2.65	1.91	5.24	3.46	2.49
	ARL →	3.44	1.80	1.17	4.24	2.22	1.45	5.29	2.77	1.81	6.09	3.19	2.08	7.94	4.16	2.72
(1,5)	d →	1.95	1.27	0.90	2.40	1.56	1.12	3.00	1.95	1.39	3.45	2.24	1.61	4.50	2.93	2.09
	ARL →	8.28	1.18	0.77	2.81	1.46	0.95	3.55	1.82	1.19	4.04	2.10	1.37	5.26	2.73	1.79

Table 1.3. Value of d and ARL when $(e_1, e_2) = (0.03, 0.1)$

(P'_0, P'_1)		α														
		0.05			0.025			0.01			0.005			0.001		
		k=1	k=2	k=3	k=1	k=2	k=3	k=1	k=2	k=3	k=1	k=2	k=3	k=1	k=2	k=3
(1, 2)	d →	4.64	3.28	2.45	5.72	4.04	3.02	7.14	5.05	3.77	8.21	5.81	4.34	10.71	7.57	5.66
	ARL →	20.69	11.22	7.33	25.48	13.81	9.02	31.82	17.24	11.26	36.60	19.84	12.96	47.73	25.87	16.90
(1, 3)		2.95	2.02	1.48	3.64	2.49	1.83	4.54	3.11	2.28	5.23	3.58	2.63	6.81	4.67	3.42
		6.97	3.71	2.42	8.58	4.57	2.98	10.71	5.70	33.72	12.33	6.56	4.28	16.07	8.55	5.58
(1, 4)		2.36	1.58	1.14	2.90	1.94	1.41	3.62	2.43	1.76	4.17	2.80	2.03	5.44	3.65	2.64
		3.90	2.05	1.34	4.80	2.53	1.65	6.00	3.15	2.06	6.90	3.63	2.37	9.00	4.73	3.09
(1,5)		2.04	1.34	0.97	2.51	1.66	1.19	3.13	2.07	1.49	3.61	2.38	1.71	4.70	3.11	2.23
		2.64	1.38	0.90	3.25	1.70	1.11	4.06	2.12	1.38	4.67	2.44	1.59	6.09	3.18	2.08

Table 1.4. Value of d and ARL when $(e_1, e_2) = (0.03, 0.30)$

(P'_0, P'_1)		α														
		0.05			0.025			0.01			0.005			0.001		
		k=1	k=2	k=3	k=1	k=2	k=3	k=1	k=2	k=3	k=1	k=2	k=3	k=1	k=2	k=3
(1, 2)	d →	5.66	3.70	2.64	6.98	4.56	3.25	8.71	5.69	4.06	10.02	6.54	4.68	13.07	8.53	6.10
	ARL →	26.55	13.96	9.11	32.69	17.20	11.22	40.81	21.47	14.01	46.95	24.70	16.12	61.22	32.21	21.01
(1, 3)		3.53	2.24	1.58	4.34	2.77	1.95	5.42	3.45	2.44	6.24	3.97	2.81	8.14	5.18	3.66
		8.88	4.62	3.02	10.94	5.69	3.72	13.66	7.10	4.64	15.71	8.17	5.34	20.49	10.65	6.96
(1, 4)		2.78	1.74	1.22	3.43	2.15	1.50	4.28	2.68	1.88	4.92	3.08	2.16	6.42	4.02	2.82
		4.97	2.57	1.68	6.12	3.16	2.07	7.65	3.95	2.58	8.80	4.54	2.97	11.47	5.92	3.88
(1,5)		2.39	1.48	1.03	2.94	1.82	1.27	3.67	2.27	1.59	4.23	2.62	1.82	5.51	3.41	2.38
		3.38	1.73	1.14	4.16	2.13	1.40	5.19	2.67	1.75	5.97	3.07	2.01	7.79	4.00	2.62

Table 1.5. Value of d and ARL when $(e_1, e_2) = (0.005, 0.02)$

(P'_0, P'_1)		α														
		0.05			0.025			0.01			0.005			0.001		
		<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>
(1, 2)	d →	4.41	3.10	2.31	5.43	3.82	2.85	6.78	4.77	3.56	7.80	5.49	4.09	10.16	7.15	5.33
	ARL →	18.22	9.85	6.43	22.44	12.13	7.92	28.01	15.15	9.89	32.23	17.43	11.38	42.02	22.72	14.84
(1, 3)		2.79	1.90	1.39	3.43	2.33	1.71	4.28	2.91	2.13	4.93	3.35	2.45	6.42	4.37	3.20
		5.97	3.17	2.07	7.35	3.90	2.55	9.18	4.87	3.18	10.56	5.60	3.66	13.77	7.31	4.77
(1, 4)		2.21	1.47	1.06	2.72	1.81	1.31	3.40	2.26	1.63	3.91	2.60	1.88	5.09	3.39	2.45
		3.26	1.71	1.12	4.01	2.10	1.38	5.01	2.63	1.72	5.76	3.02	1.97	7.51	3.94	2.57
(1,5)		1.90	1.24	0.89	2.34	1.53	1.09	2.92	1.91	1.37	3.36	2.19	1.57	4.38	2.86	2.05
		2.15	1.12	0.73	2.65	1.38	0.90	3.31	1.72	1.13	3.80	1.98	1.30	4.96	2.58	1.69

Table 1.6. Values of ϕ for different values of (e_1, e_2)

(P'_0, P'_1) → (e_1, e_2) ↓		(1, 2)			(1, 3)			(1, 4)			(1, 5)		
		<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>	<i>k=1</i>	<i>k=2</i>	<i>k=3</i>
(0, 0)		67.50	73.70	77.68	69.78	75.90	79.58	71.31	77.30	80.75	72.45	78.31	81.57
(0.01, 0.05)		68.24	74.42	78.31	70.40	76.47	80.06	71.85	77.77	81.14	72.92	78.71	81.89
(0.03, 0.1)		67.45	73.64	77.63	69.54	75.67	79.39	70.92	76.95	80.46	71.94	77.86	81.21
(0.03, 0.30)		72.96	78.70	81.88	74.61	80.07	82.96	75.69	80.94	83.63	76.48	81.55	84.09
(0.005, 0.02)		67.76	73.96	77.91	69.99	76.10	79.75	71.49	77.46	80.88	72.60	78.44	81.67

Table- 1.6 shows that the angle of mask, ϕ increases as the value of k increases. For comparison, if we take $(e_1, e_2) = (0,0), (0.005, 0.02), (0.03,0.3), (0.01, 0.05)$, the angle of mask is 67.50, 67.76, 68.24, and 72.96 respectively, which is the increased value of the mask, whereas $(e_1, e_2) = (0.03,0.1)$ the angle of mask is 67.45, lesser as compared to error free case.

Competing Interests

Authors have declared that no competing interests exist.

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